

Covariance pattern in LMMstar

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1 LV pattern

1.1 Theory

Consider 4 timepoints. The traditional parametrisation of the residual variance-covariance matrix of a factor model is

$$\Omega = \begin{bmatrix} \omega_1^2 + \tau & \cdot & \cdot & \cdot \\ \lambda_2\tau & \omega_2^2 + \lambda_2^2\tau & \cdot & \cdot \\ \lambda_3\tau & \lambda_2\lambda_3\tau & \omega_3^2 + \lambda_3^2\tau & \cdot \\ \lambda_4\tau & \lambda_2\lambda_4\tau & \lambda_3\lambda_4\tau & \omega_4^2 + \lambda_4^2\tau \end{bmatrix}$$

LMMstar uses a different parametrisation with distinct parameters for the variance and the correlation:

$$\Omega = \begin{bmatrix} \sigma_1^2 & \cdot & \cdot & \cdot \\ \rho_1\rho_2\sigma_1\sigma_2 & \sigma_2^2 & \cdot & \cdot \\ \rho_1\rho_3\sigma_1\sigma_3 & \rho_2\rho_3\sigma_2\sigma_3 & \sigma_3^2 & \cdot \\ \rho_1\rho_4\sigma_1\sigma_4 & \rho_2\rho_4\sigma_2\sigma_4 & \rho_3\rho_4\sigma_3\sigma_4 & \sigma_4^2 \end{bmatrix}$$

The two parametrisation are equivalent when assuming the same sign for all correlation (e.g. all positive correlation).

$\omega, \tau, \lambda \rightarrow \rho, \sigma$: the σ values can be deduce from the diagonal of Ω .

To get the ρ value, we can first multiply $\rho_1\rho_2\sigma_1\sigma_2 = \lambda_2\tau$ by $\rho_1\rho_3\sigma_1\sigma_3 = \lambda_3\tau$ and divide by $\rho_2\rho_3\sigma_2\sigma_3 = \lambda_2\lambda_3\tau$:

$$\begin{aligned} \frac{\rho_1^2\rho_2\rho_3\sigma_1^2\sigma_2\sigma_3}{\rho_2\rho_3\sigma_2\sigma_3} &= \frac{\lambda_2\lambda_3\tau^2}{\lambda_2\lambda_3\tau} \\ \rho_1^2\sigma_1^2 &= \tau \\ \rho_1^2 &= \frac{\tau}{\omega_1^2 + \tau} \end{aligned}$$

More generally, denoting $\lambda_1 = 1$, from:

$$\begin{aligned}\sigma_i^2 &= \omega_i^2 + \lambda_i^2 \tau \\ \rho_i \rho_j \sigma_i \sigma_j &= \lambda_i \lambda_j \tau\end{aligned}$$

we can deduce

$$\begin{aligned}\rho_i^2 \rho_j \rho_k \sigma_i^2 \sigma_j \sigma_k &= \lambda_i^2 \lambda_j \lambda_k \tau^2 \\ \rho_i^2 \sigma_i^2 &= \lambda_i^2 \tau \\ \rho_i &= \sqrt{\frac{\lambda_i^2 \tau}{\omega_i^2 + \lambda_i^2 \tau}}\end{aligned}$$

Technical ρ_i could be negative but here use the assumption of same (positive) sign for all correlations.

$\boldsymbol{\rho}, \boldsymbol{\sigma} \rightarrow \boldsymbol{\omega}, \boldsymbol{\tau}, \boldsymbol{\lambda}$: we can re-use the previous result:

$$\rho_i^2 = \frac{\lambda_i^2 \tau}{\omega_i^2 + \lambda_i^2 \tau} = \frac{\lambda_i^2 \tau}{\sigma_i^2}$$

So for $i = 1$:

$$\tau = \rho_1^2 \sigma_1^2$$

and otherwise:

$$\begin{aligned}\lambda_i^2 &= \frac{\rho_i^2 \sigma_i^2}{\rho_1^2 \sigma_1^2} \\ \lambda_i &= \text{sign}(\rho_1 \rho_i) \frac{\rho_i \sigma_i}{\rho_1 \sigma_1}\end{aligned}$$

One can then deduce ω_i :

$$\omega_i = \sqrt{\sigma_i^2 - \lambda_i \tau} = \sigma_i \sqrt{1 - \rho_i^2}$$

1.2 Example

Simulate data

```
library(lava)
library(LMMstar)

mSim <- lvm(c(Y1,Y2,Y3,Y4)~eta+age)
latent(mSim) <- ~eta

set.seed(10)
n <- 100
n.time <- length(endogenous(mSim))
dfW.sim <- cbind(id = paste0("Id",1:n), sim(mSim, n = n, latent = FALSE))
dfW.sim$id <- factor(dfW.sim$id, unique(dfW.sim$id))
head(dfW.sim)
```

	id	Y1	Y2	Y3	Y4	age
1	Id1	0.8087642	0.02821369	2.0055318	2.29256267	0.8694750
2	Id2	0.3174894	0.92111736	0.8326184	1.09215142	-0.6800096
3	Id3	0.9880281	1.31941524	3.7496337	1.72867315	0.1732145
4	Id4	-0.3524308	0.95831086	1.1187839	1.03908643	-0.1594380
5	Id5	0.3496855	-0.57807269	-1.0256767	0.18052490	0.7934994
6	Id6	0.1276581	0.30103845	0.2336854	0.06061876	1.6943505

Convert to long format:

```
dfL.sim <- reshape(dfW.sim, direction = "long", idvar = c("id","age"),
  varying = paste0("Y",1:4), sep="")
dfL.sim$time <- as.factor(dfL.sim$time)
rownames(dfL.sim) <- NULL
head(dfL.sim)
```

	id	age	time	Y
1	Id1	0.8694750	1	0.8087642
2	Id2	-0.6800096	1	0.3174894
3	Id3	0.1732145	1	0.9880281
4	Id4	-0.1594380	1	-0.3524308
5	Id5	0.7934994	1	0.3496855
6	Id6	1.6943505	1	0.1276581

Fit LVM:

```
m.lvm <- lvm(c(Y1,Y2,Y3,Y4)~eta+age, eta ~ 0)
latent(m.lvm) <- ~eta
e.lvm <- estimate(m.lvm, data = dfW.sim)
logLik(e.lvm)
```

'log Lik.' -651.0478 (df=16)

Export coefficient by type:

```
mu.lvm <- coef(e.lvm)[c(paste0("Y",1:n.time),paste0("Y",1:n.time,"~age"))]  
lambda.lvm <- c(1,coef(e.lvm)[paste0("Y",2:n.time,"~eta")])  
tau.lvm <- coef(e.lvm)["eta~~eta"]  
omega.lvm <- coef(e.lvm)[paste0("Y",1:n.time,"~~Y",1:n.time)]  
list(mu = mu.lvm,  
      lambda = lambda.lvm,  
      tau = tau.lvm,  
      omega = omega.lvm)
```

\$mu

	Y1	Y2	Y3	Y4	Y1~age	Y2~age	Y3~age
	-0.1835368	-0.1491306	-0.0194078	0.1459640	0.9502774	1.0535363	0.9671297
Y4~age							
	1.0349377						

\$lambda

	Y2~eta	Y3~eta	Y4~eta
	1.0000000	0.8653753	1.1024519
		1.0537868	

\$tau

eta~~eta
1.259395

\$omega

	Y1~~Y1	Y2~~Y2	Y3~~Y3	Y4~~Y4
	0.8444582	0.9968492	1.0012298	0.9868099

Conversion to LMM coefficients:

```
list(sigma = sqrt(omega.lvm + lambda.lvm^2 * tau.lvm),  
      rho = sqrt( lambda.lvm^2 * tau.lvm / (omega.lvm + lambda.lvm^2 * tau.  
lvm)))
```

\$sigma

	Y1~~Y1	Y2~~Y2	Y3~~Y3	Y4~~Y4
	1.450466	1.392831	1.591194	1.544450

\$rho

	Y2~eta	Y3~eta	Y4~eta
	0.7737012	0.6972477	0.7775305
		0.7657021	

Fit LMM:

```
rhoLVM <- function(p,time,...){
  R <- tcrossprod(p[time])
  diag(R) <- 1
  return(R)
}
myStruct <- CUSTOM(~time,
  FCT.sigma = function(p,time,X){p[time]},
  init.sigma = setNames(rep(1.45,n.time),paste0("sigma",1:n.time)),
  FCT.rho = rhoLVM,
  init.rho = setNames(rep(0.7,n.time),paste0("rho",1:n.time)))
e.lmmCUSTOM <- lmm(Y ~ time*age,
  repetition = ~time|id,
  structure = myStruct, data = dfL.sim,
  method.fit = "ML")
logLik(e.lmmCUSTOM)
```

```
[1] -651.0478
```

We get exactly the same log-likelihood as the latent variable model. Export coefficient by type:

```
mu.lmm <- coef(e.lmmCUSTOM, effects = "mean")
sigma.lmm <- coef(e.lmmCUSTOM, effects = "variance")
rho.lmm <- coef(e.lmmCUSTOM, effects = "correlation")
list(mu = mu.lmm,
  sigma = sigma.lmm,
  rho = rho.lmm)
```

```
$mu
(Intercept)      time2      time3      time4      age  time2:age
-0.18353676  0.03440620  0.16412896  0.32950080  0.95027744  0.10325886
  time3:age  time4:age
  0.01685224  0.08466026
```

```
$sigma
  sigma1  sigma2  sigma3  sigma4
1.450466 1.392831 1.591194 1.544450
```

```
$rho
  rho1  rho2  rho3  rho4
0.7737013 0.6972475 0.7775306 0.7657020
```

Conversion to LVM coefficients:

```
list(lambda = rho.lmm*sigma.lmm/(rho.lmm[1]*sigma.lmm[1]),  
      tau = rho.lmm[1]^2*sigma.lmm[1]^2,  
      omega = sigma.lmm^2 * (1-rho.lmm^2))
```

\$lambda

	rho1	rho2	rho3	rho4
	1.000000	0.865375	1.102452	1.053786

\$tau

	rho1
	1.259395

\$omega

	sigma1	sigma2	sigma3	sigma4
	0.8444578	0.9968496	1.0012293	0.9868104